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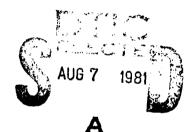
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# TECHNICAL REPORT ARBRL-TR-02331

# QUANTILES OF MULTI-SAMPLE SMIRNOV TYPE STATISTICS

Mark P. Becker Malcolm S. Taylor

June 1981



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# US ARMY ARMAMENT RESEARCH AND DEVELOPMENT COMMAND BALLISTIC RESEARCH LABORATORY ABERDEEN PROVING GROUND, MARYLAND

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Let $X_1^{(1)}, X_2^{(i)}, \dots, X_{n_i}^{(i)}$ , $i = 1, 2, \dots, c$ , be samp	oles of c independent random
variables $X^{(i)}$ with continuous cdf $F^{(i)}$ . To test	the hypothesis $H_a$ : $\Gamma^{(i)}$ =
$F^{(j)}$ , i, j = 1,2,, c, Birnbaum and Hall introdu	iced the statistics
$D(n_1, n_2, \ldots, n_c)$ and $D^{T}(n_1, n_2, \ldots, n_c)$ and deter	mined their exact small
sample distributions for $c = 2,3$ and $n_i = n_j$ . Thi	s paper considers exact small
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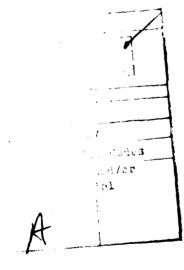
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#### I. INTRODUCTION AND SUMMARY

Let  $X_1^{(i)}, X_2^{(i)}, \ldots, X_{n_i}^{(i)}$ ,  $i = 1, 2, \ldots$ , c be samples of c independent random variables  $X^{(i)}$  with continuous cumulative distribution functions  $F^{(i)}$  and empirical distributions  $F^{*(i)}$ 

$$F^{*(i)}(x) = 0$$
  $x < X_1^{(i)}$   
 $= k/n_i$   $X_k^{(i)} \le x < X_{k+1}^{(i)}$ ,  $1 \le k < n_i$   
 $= 1$   $X_{n_i}^{(i)} \le x$ .

We assume, without loss of generality, that  $X_1^{(i)} < X_2^{(i)} < \ldots < X_{n_i}^{(i)}$ .

To test the hypothesis that the population distribution functions are identical, Birnbaum and Hall introduced the statistics

$$D(n_1, n_2, ..., n_c) = \sup_{x,i,j} |F^{*(i)}(x) - F^{*(j)}(x)|,$$
 (1)

$$D^{+}(n_{1},n_{2}, ...,n_{c}) = \sup_{x,i < j} [F^{+(i)}(x) - F^{+(j)}(x)]$$
 (2)

i,j = 1,2, ..., c which are generalizations of the well known Kolmogorov-Smirnov two sample statistics D(m, n) and  $D^{+}(m, n)$ .

Under the null hypothesis

$$H_0: F^{(i)} = F^{(j)}$$
 i, j = 1,2, ..., c,

the exact small sample distributions of the statistics (1) and (2) were determined by Birnbaum-Hall using difference equations that could be solved recursively and tabled values of

$$P[D(n, n, n) \le r], P[D(n,n) \le r], P[D^+(n,n) \le r]$$

were developed for selected values of n between 1 and 40 and of r = k/n,  $k = 1, 2, \ldots, n$ .

<sup>&</sup>lt;sup>1</sup>Z. W. Birnbaum and R. A. Hall, "Small Sample Distributions for Multi-Sample Statistics of the Smirnov Type," Ann. Math. Stat., Vol. 31 (1960), pp. 710-720.

The limitations on the tables were for the most part imposed by the speed and capacity of the computer necessary to solve the difference equations. This paper utilizes some notions from graph theory, and in so doing, provides exact small sample distributions for more than three samples and for unequal sample sizes. Representative tables are included.

#### II. A GEOMETRICAL PERSPECTIVE

For a geometrical perspective of the problem we parallel in the next two paragraphs the development of Birnbaum-Hall.

The c-dimensional random variable  $[F^{*(1)}(x), F^{*(2)}(x), \ldots, F^{*(c)}(x)]$ , for fixed x, takes on values  $(k_1/n_1, k_2/n_2, \ldots, k_c/n_c), k_i \in \{0, 1, \ldots, n_i\}$ ;  $i = 1, 2, \ldots, c$ , which lie in the unit hypercube. The transformation  $(x_1, x_2, \ldots, x_c)$  into  $(y_1, y_2, \ldots, y_c)$  defined by  $y_i = n_i x_i$  maps the hypercube into a c-dimensional hypervolume with sides  $(n_1, n_2, \ldots, n_c)$  and the random c-vector into points  $(k_1, k_2, \ldots, k_c)$  with integer valued coordinates.

Under the null hypothesis  $H_0$ :  $F^{(i)} = F^{(j)}$  i,  $j = 1, 2, \ldots$ , c we may consider the c samples as coming from the same population with N! equally likely ways of drawing the combined sample of size  $N = n_1 + n_2 + \ldots + n_c$ . Now suppose the values of the combined ordered sample are located on the real line, and the null c-vector is used as a counter in the following manner: Starting at a value less than min  $\{\chi_1^{(1)}, \chi_1^{(2)}, \ldots, \chi_1^{(c)}\}$  with the counter equal  $(0, 0, \ldots, 0)$  and moving in the direction of increasing magnitude, the kth coordinate of the counter is incremented by one whenever a sample value  $X_i^{(k)}$  is encountered. In so doing, a 1-1 correspondence between a path in c-space from  $(0,0,\ldots,0)$  to  $(n_1,n_2,\ldots,n_c)$  and the c samples is defined.

The total number of such equally likely paths is the multinomial coefficient N!/ $(n_1!n_2! \dots n_c!)$ . The critical region for the test will consist of a set of points R =  $\{(k_1, k_2, \dots, k_c)\}$  such that any path from  $(0,0,\dots,0)$  to  $(n_1,n_2,\dots,n_c)$  passing through R will cause  $H_0$  to be rejected. If we denote the number of paths passing through R as  $Q(n_1,n_2,\dots,n_c;R)$  then the probability of an error of the first kind

will be

$$\alpha = \frac{Q(n_1, n_2, ..., n_c; R)}{N!/(n_1!n_2! ... n_c!)}.$$

#### III. DETERMINATION OF EXACT SMALL SAMPLE DISTRIBUTION

Determination of the exact small sample distribution of (1) and (2) from this approach involves an enumerative process to determine  $Q(n_1,n_2,\ldots,n_c;R)$ . The set R defining the critical region suggests itself as  $R = \{(k_1,k_2,\ldots,k_c) \mid \sup_{i,j} \mid n_j k_i - n_i k_j \mid n_j n_j r\}$  and

$$R^+ = \{(k_1, k_2, ..., k_c) \mid \sup_{i < j} (n_j k_i - n_i k_j) > n_i n_j r\}$$
 corresponding to

 $D(n_1, n_2, \ldots, n_c) > r$  and  $D^+(n_1, n_2, \ldots, n_c) > r$  respectively. These regions of rejection are again extensions of the Kolmogorov-Smirnov procedure; however, there is no theoretical basis for their limitation to this form.

Considering the paths in c-space from  $(0,0,\ldots,0)$  to  $(n_1,n_2,\ldots,n_c)$  as simple paths in a finite directed graph, (see e.g. Berge<sup>2</sup>) we have at our disposal a method of enumeration appearing in a paper by Pototchi<sup>3</sup> which seems well suited to computer manipulation. This method, while conceptually analogous to the difference equation evaluated by Birnbaum-Hall, allows us to accomplish the enumeration necessary to evaluate  $Q(n_1,n_2,\ldots,n_c;R)$  and to consider the number of samples c>3 as well as unequal sample sizes in an efficient and economical manner.

#### IV. TABLES

Table 1 describes the upper tail of the distribution of D, and contains the probabilities  $P[D(n_1,n_2,n_3) \le r]$  for  $5 \le n_1,n_2,n_3 \le 10$  starting at the largest increment of r = .1(.1)1.0 for which the tail probability equals or exceeds .20.

<sup>&</sup>lt;sup>2</sup>C. Berge, The Theory of Graphs and its Applications, John Wiley & Sons, New York, 1962.

A. Pototchi, "A Simple Algorithm for Determining the Number of Paths in a Finite Graph," Economic Cyber Studies Res., Vol. 2 (1967), pp. 81-85.

Table 2 contains the probabilities  $P[D(n,n,n,n) \le r]$  for n = 2(1)10 corresponding to values of r = k/n consistent with the differences  $|F^{*(i)}(x) - F^{*(j)}(x)|$  which can occur.

The values in Table 1 for the triples (n,n,n), n=5(1)10 correspond to those computed by Birnbaum-Hall. The values in Table 2 for the 4-tuples (n,n,n,n), n=5,10 are consistent with those Monte Carloed by Gardner, 4 et al. Taken together, these provide additional verification of the accuracy of the tables.

<sup>&</sup>lt;sup>4</sup>R. H. Gardner, J. E. Pinder III, and R. S. Wood, "Monte Carlo Estimation of Percentiles for the Multi-Smirnov Test," J. Statist. Comput. Simul., Vol. 10 (1980), pp. 243-249.

TABLE 1. QUANTILES OF THE BIRNBAUM-HALL, OR THREE SAMPLE SMIRNOV, STATISTIC FOR UNEQUAL SAMPLE SIZES.

<sup>n</sup> 1	n <sub>2</sub>	n <sub>3</sub>	r	$P[D(n_1,n_2,n_3) \leq r]$	n <sub>1</sub>	n <sub>2</sub>	n <sub>3</sub>	r	$P[D(n_1,n_2,n_3) \leq r]$
5	5	5	.6 .7 .8 .9	.81109 .81109 .97819 .97819 1.00000	5	6	6	.6 .7 .8 .9	.64253 .89712 .93216 .98982 1.00000
5	5	6	.6 .7 .8 .9	.67353 .85609 .94628 .98457	5	6	7	.6 .7 .8 .9	.70972 .87336 .95008 .99238 1.00000
5	5	7	.6 .7 .8 .9	.75909 .83058 .96542 .98776 1.00000	5	6	8	.6 .7 .8 .9	.70123 .88876 .95968 .99372 1.00000
5	5	8	.6 .7 .8 .9	.74472 .86485 .97582 .98947 1.00000	5	6	9	.6 .7 .8 .9	.72578 .91119 .96508 .99446 1.00000
5	5	9	.5 .6 .7 .8 .9	.46916 .80259 .88479 .98176 .99043	5	6	10	.6 .7 .8 .9	.75753 .92490 .96825 .99490 1.00000
5	5	10	.5 .6 .7 .8 .9	.52097 .83928 .89685 .98531 .99100 1.00000	5	7	7	.6 .7 .8 .9	.77668 .84966 .96547 .99462 1.00000
					5	7	8	.6 .7 .8 .9	.74794 .88248 .97350 .99578 1.00000

TABLE 1. (CONT'D)

$n_1$	n <sub>2</sub>	n <sub>3</sub>	r	$P[D(n_1,n_2,n_3) \leq r]$	<u>n</u> 1	n <sub>2</sub>	n <sub>3</sub>	r	$P[D(n_1,n_2,n_3) \leq r]$
5	7	9	.6 .7 .8 .9	.78143 .90138 .97793 .99642 1.00000	5	10	10	.5 .6 .7 .8 .9	.70953 .88851 .96359 .99235 .99870 1.00000
5	7	10	.5 .6 .7 .8 .9	.54453 .81483 .91269 .98047 .99680 1.00000	6	6	6	.6 .7 .8 .9	.68408 .93216 .93216 .99383 1.00000
5	8	8	.6 .7 .8 .9	.74256 .91251 .98065 .99685 1.00000	6	6	7	.6 .7 .8 .9	.74897 .91138 .94990 .99569 1.00000
5	8	9	.6 .7 .8 .9	.78949 .92949 .98457 .99745 1.00000	6	6	8	.6 .7 .8 .9	.74664 .90488 .95945 .99662 1.00000
5	8	10	.5 .6 .7 .8 .9	.62530 .81808 .93951 .98681 .99780 1.00000	6	6	9	.6 .7 .8 .9	.74555 .92900 .96485 .99712 1.00000
5	9	9	.5 .6 .7 .8 .9	.59437 .83481 .94532 .98821 .99803	6	6	10	.6 .7 .8 .9 1.0	.78119 .94384 .96806 .99739 1.00000
5	9	10	.5 .6 .7 .8 .9	.65118 .86208 .95464 .99031 .99836 1.00000	6	7	7	.5 .6 .7 .8 .9	.49330 .81149 .88975 .96517 .99719

TABLE 1. (CONT'D)

<u>n</u> 1	n <sub>2</sub>	n <sub>3</sub>		$P[D(n_1, n_2, n_3) \leq r]$	<sup>n</sup> 1	n <sub>2</sub>	n <sub>3</sub>	r	$P[D(n_1, n_2, n_3) \leq r]$
6	7	8	.6	.78637	6	9	10	.5	.67974
			.7	.90257				.6	.84898
			.8	.97319				.7	.95453
			.9	.99792				.8	.99011
			1.0	1.00000				.9	.99934
								1.0	1.00000
6	7	9	.6	.79154	6	10	10	.5	.70263
			.7	.92361				.6	.87756
			.8	.97765				. 7	.96382
			. 9	.99830				. 5	.99217
			1.0	1.00000				9	.99950
			1.0	1.0000				1.0	1.00000
6	7	10	.5	.57700	7	7	7	.5	.56209
•	•		.6	.82881	•	•	,	.6	.86823
			.7	.93638				.7	.86823
			.8	.98023				.8	.97750
			.9	.99851					
								.9	.99830
			1.0	1.00000				1.0	1.00000
6	8	8	.6	.78567	7	7	8	.5	.63985
			.7	.91100				. 6	.81671
			.8	.98035				.7	.89916
			.9	.99855				.8	.98371
			1.0	1.00000				.9	.99883
								1.0	1.00000
6	8	9	.5	.61580	7	7	9	.5	.61004
			.6	.80505				.6	.82936
			.7	.92848				.7	.91681
			.8	.98430				. 8	.98702
			.9	.99888				.9	.99909
			1.0	1.00000				1.0	1.00000
6	8	10	.5	.65020	7	7	10	.5	.62594
			.6	.83826				.6	.86663
			.7	.93879				.7	.92731
			.8	.98659				.8	.98888
			.9	.99906				.9	.99922
			1.0	1.00000				1.0	1.00000
6	9	9	.5	.65335	7	8	8	.6	.79392
			.6	.81993				. 7	.92685
			.7	.94488				.8	.98888
			.8	.98798				.9	.99924
			.9	.99918				1.0	1.00000
			1.0	1.00000				1.0	1.00000
			1.0	1.00000					

TABLE 1. (CONT'D)

n <sub>1</sub>	n <sub>2</sub>	n <sub>3</sub>	r	$P[D(n_1, n_2, n_3) \le r]$	n <sub>1</sub>	n <sub>2</sub>	n <sub>3</sub>	r	$P[D(n_1,n_2,n_3) \leq r]$
7	8	9	.5	.64705	8	8	10	.5	.74735
			.6	.82084		-		.6	.86362
			.7	.94230				.7	.96997
			.8	.99158				.8	.99601
			.9	.99944				.9	.99976
			1.0	1.00000				1.0	1.00000
			1.0	1.00000				1.0	1.00000
7	8	10	.5	.68971	8	9	9	.5	.65584
			.6	.85147	_	_	_	.6	.87789
			.7	.95131				.7	.97370
			.8	.99306				.8	.99659
			.9	.99955				.9	.99980
			1.0	1.00000				1.0	1.00000
			1.0	1.00000				1.0	1.00000
7	9	9	.5	.66124	8	9	10	.5	.71259
			.6	.84333				.6	.90160
			.7	.95640				.7	.97966
			.8	.99394				.8	.99745
			.9	.99962				.9	.99985
			1.0	1.00000				1.0	1.00000
7	9	10	.5	.69235	8	10	10	.5	.76924
			.6	.87042				.6	.92368
			.7	.96455				.7	.98495
			.8	.99523				.8	.99819
			.9	.99970				.9	.99990
			1.0	1.00000				1.0	1.00000
_			_						
7	10	10	.5	.72037	9	9	9	.5	.71542
			.6	.89658				.6	.91350
			.7	.97223				.7	.98248
			.8	.99641				.8	.99785
			.9	.99979				.9	.99988
			1.0	1.00000				1.0	1.00000
8	8	8	.6	.79392	9	9	10	.5	.76978
		-	.7	.95029	-		10	.6	
			.8	.99292				.7	.93369 .98714
			.9	.99954				.8	.99848
			1.0	1.00000				.9	
			1.0	1.00000				1.0	.99992 1.00000
								1.0	1.00000
8	8	9	.5	.67466	9	10	10	. 4	.53864
			.6	.83749				.5	.82180
			.7	.96289				.6	.95150
			.8	.99494				.7	.99104
			.9	.99968				.8	.99898
			1.0	1.00000				.9	.99995
								1.0	1.00000
								1.0	1.0000

TABLE 1. (CONT'D)

n <sub>1</sub>	n <sub>2</sub>	n <sub>3</sub>	r	$P[D(n_1,n_2,n_3) \leq r]$
10	10	10	. 4	.63715
			.5	.86930
			.6	.96645
			.7	.99411
			.8	.99936
			.9	.99997
			1.0	1 00000

TABLE 2. QUANTILES OF A FOUR SAMPLE SMIRNOV STATISTIC FOR EQUAL SAMPLE SIZES.

n <sub>1</sub>	n <sub>2</sub>	n <sub>3</sub>	n <sub>4</sub>	r	$P[D(n_1, n_2, n_3, n_4) \le r]$	n,	n a	2 n	3 <sup>n</sup> 4	r	$P[D(n_1, n_2, n_3, n_4) \le r]$
2	2	2	2	.5 1.0	.22857	8	8	8	8	.500	
				***	1.00000					.750	
										.875	
										1.000	
3	3	3	3	.333		9	9	9	9	.444	.56199
				. 666	6 .64286					.555	
				1.000	1.00000					.666	
										.777	
										.888	
										1.000	
4	4	4	4	.25	.00526	10	10	10	10	.5	.78000
				.50	.35707					.6	.93879
				.75	.87219					.7	.98874
				1.00	1.00000					.8	.99875
										.9	.99994
										1.0	1.00000
5	5	5	5	.2	.00068						
				. 4	.18184						
				.6	.69490						
				.8	.95982						
				1.0	1.00000						
6	6	6	6	.500							
				.666							
				.833	.98824						
				1.000	1.00000						
7	7	7	7	<b>57</b> 1	77075						
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